

等価重み粒子フィルタによる時変パラメータの推定

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Introduction

Motivation: To improve robustness and efficiency for estimating the state and time-varying parameters in nonlinear high-dimensional systems.

Difficulty: Typical state augmentation technique tend to fail to detect parameter changes unless there are strong enough correlations.

Solution: Introduce correlated perturbation and use an efficient particle filter "IEWPF" [1] combined with adaptive moment estimation (Adam [2]) based optimization technique from machine learning.

Methods

State augmentation with correlated perturbation

Time evolution of state in augmented state-space model:

$$z^n = \tilde{f}(z^{n-1}) + \tilde{\beta}^n = \begin{pmatrix} f(x^{n-1}, \theta^n) \\ \theta^{n-1} \end{pmatrix} + \begin{pmatrix} \beta^n \\ \eta^n \end{pmatrix}$$

$$y^n = H(z^n) + \varepsilon^n$$

can approximately express by using Taylor series expansion:

$$f(x^{n-1}, \theta^n) = f(x^{n-1}, \theta^{n-1}) + \frac{\partial f}{\partial \theta} (\theta^n - \theta^{n-1}).$$

Hence,

$$z^n = \begin{pmatrix} x^n \\ \theta^n \end{pmatrix} = \begin{pmatrix} f(x^{n-1}, \theta^{n-1}) \\ \theta^{n-1} \end{pmatrix} + \tilde{\beta}^n, \text{ where } p(\tilde{\beta}^n) = N(0, Q_{aug}^n),$$

$$Q_{aug}^n = \begin{pmatrix} \text{cov}[x^n, x^n] & \text{cov}[x^n, \theta^n] \\ \text{cov}[\theta^n, x^n] & \text{cov}[\theta^n, \theta^n] \end{pmatrix} = \begin{pmatrix} Q_\beta + \frac{\partial f}{\partial \theta} Q_\eta \frac{\partial f^T}{\partial \theta} & \frac{\partial f}{\partial \theta} Q_\eta \\ \left(\frac{\partial f}{\partial \theta} Q_\eta\right)^T & Q_\eta \end{pmatrix}.$$

Combination of IEWPF and parameter optimization

- The Implicit Equal-Weights Particle Filter (IEWPF) [1] can avoid filter degeneracy by using a proposal density which the weight of each particle can be equal to the target weight.
- Assume typical geophysical systems which time step interval T exist between observations, and introduce new proposal distribution during this period as shown in Fig. 1.

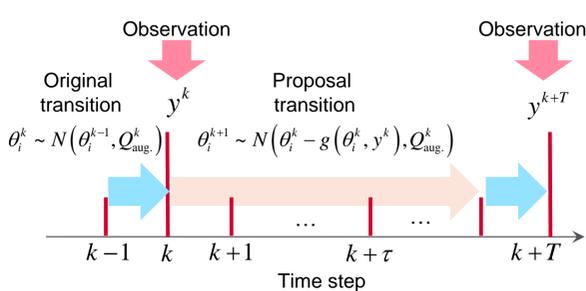


Fig.1. Parameter updating diagram from one step before observed step k to next observation $k+T$. Function g denotes parameter correction term calculated by using the last observation and τ is the number of time steps after the observation.

Adam-based parameter correction term

- Using gradient of log-likelihood at the last observation step, update momentum m and norm (squared gradient) v as following equation. μ and ρ denote hyper-parameter that control these moving average.

$$\nabla_\theta L_t^k \equiv -2\nabla_\theta \ln[p(y^k | x_t^{k-1}, \theta_t^{k-1})],$$

$$\text{where } p(y^k | x_t^{k-1}, \theta_t^{k-1}) \propto \exp\left[-\frac{1}{2}(y^k - Hf(x_t^{k-1}, \theta_t^{k-1}))^T R^{-1}(y^k - Hf(x_t^{k-1}, \theta_t^{k-1}))\right]$$

$$m_t^k = \mu m_t^{k-1} + (1-\mu)\nabla_\theta L_t^k, \quad v_t^k = \rho v_t^{k-1} + (1-\rho)(\nabla_\theta L_t^k)^2$$

- Then parameter correction term g is expressed as follows. λ_0 and δ denote step-size and a term to avoid division by 0, respectively.

$$g(\theta_t^k, y^k) = \frac{\lambda_0}{\tau} \frac{\hat{m}^k}{\sqrt{\hat{v}^k + \delta}} \theta^{k-1}, \text{ where } \hat{m}^k = \frac{m^k}{1-\mu}, \quad \hat{v}^k = \frac{v^k}{1-\rho}$$

References

- Zhu M., van Leeuwen PJ, Amezcuca J., 2016, 'Implicit equal-weights particle filter', Q. J. R. Meteorol. Soc. 142: 1904-1919.
- P. Kingma D. and Lei Ba J., 'Adam: A method for stochastic optimization', conference paper at ICLR 2015

Mixture proposal density

Remaining problem has been the choice of the step-size factor. Especially in proposed method combined with Adam-based correction term, step-size plays a critical role in balancing a stable estimation with fast tracking to the abrupt changes. Therefore, in order to improve both stability and follow-up capability for uncertain environment, mixture proposal density is introduced as follows.

Particle

$$q(z_i^{n+1} | z_i^n) = \phi \mathcal{N}\left(\begin{pmatrix} f(x_i^n, \theta_i^n) \\ \theta_i^n - g(\theta_i^n, y^n, \lambda_0) \end{pmatrix}, Q_{aug}^n\right) + (1-\phi) \mathcal{N}\left(\begin{pmatrix} f(x_i^n, \theta_i^n) \\ \theta_i^n - g(\theta_i^n, y^n, c\lambda_0) \end{pmatrix}, Q_{aug}^n\right).$$

Case study: Lorenz-96 with parameterized forcing

Proposed method is tested with twin experiment using Lorenz-96 model whose forcing is parameterized as follows:

$$\frac{d}{dt} x_j = (x_{j+1} - x_{j-2})x_{j-1} - x_j + F_j, \quad F_j = f_0 + \theta_1 \sin\left(\frac{2\pi}{\theta_2} j\right).$$

x_j ($j=1, \dots, N$) is the state variables at position j of dimension N_j and f_0 , θ_1 , θ_2 is unknown or time-varying parameter. In the following experiments, the true model error covariance matrix Q is chosen as tridiagonal matrices, the main diagonal value is 0.20 and both sub- and superdiagonal values are 0.05. The observation error matrix R is diagonal with main diagonal value 0.01. For the assimilation, we choose same matrix Q and R as when observation was generated.

Evaluation results of mixture method ($N_j = 40$)

In order to investigate the ability of the proposed mixture method to estimate time-varying parameters, a trajectory comparing results of different step-sizes (a)(b) and mixture (c) is shown in Fig.2. Compared to the small step-size (a), a reduction of about 20% of the time-averaged RMSE is achieved with the proposed mixing method (c).

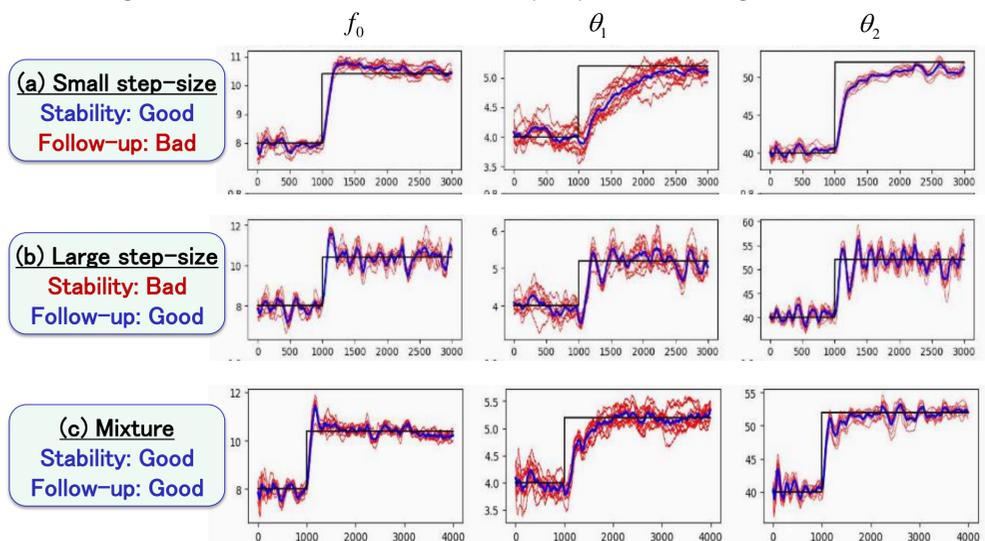


Fig.2. Time-series results of 40-dimensional Lorenz-96 with 10 ensemble members. All variables are observed by every 4th step and the observations are generated from the true parameter value [f_0 , θ_1 , θ_2] = [8.0, 4.0, 40.0] changed 30% bigger at the 1000th step.

Conclusion

Proposed method has capability to track the change of the model characteristics (e.g. parameters) by extremely fewer ensemble even in case of high-dimensional application.