# Dependence Structure of Bivariate Order Statistics and its Applications 

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## Abstract

We study the dependence structure of bivariate order statistics，and prove that if the underlying bivariate distribution $H$ is positive quadrant depen－ dent（PQD）then so is each pair of bivariate order statistics．As an applica－ tion，we show that if $H$ is PQD，the bivariate distribution $K_{+}^{(n)}$ ，proposed by Bairamov and Bayramoglu（2012），is greater than or equal to Baker＇s（2008） distribution $H_{+}^{(n)}$ ．We also show that if $H$ is PQD，$K_{+}^{(n)}$ converges weakly to the Fréchet－Hoeffding upper bound as $n$ tends to infinity．

## Introduction

## Bivariate order statistics

Suppose that $\left(X_{1}, Y_{1}\right), \ldots,\left(X_{n}, Y_{n}\right) \sim^{\text {i．i．d．}} H(x, y)=\operatorname{Pr}(X \leq x, Y \leq y)$ ．

$$
\text { Marginals : } F(x):=\operatorname{Pr}(X \leq x) ; \quad G(y):=\operatorname{Pr}(Y \leq y)
$$

Order statistics：$X_{1, n} \leq X_{2, n} \leq \cdots \leq X_{n, n} ; \quad Y_{1, n} \leq Y_{2, n} \leq \cdots \leq Y_{n, n}$
Distribution functions：

$$
\begin{aligned}
F_{r, n}(x) & :=\operatorname{Pr}\left(X_{r, n} \leq x\right)=\sum_{i=r}^{n}\binom{n}{i} F^{i}(x)(1-F(x))^{n-i} \\
G_{s, n}(y) & :=\operatorname{Pr}\left(Y_{s, n} \leq y\right)=\sum_{j=s}^{n}\binom{n}{j} G^{j}(y)(1-G(y))^{n-j}
\end{aligned}
$$

When $X$ and $Y$ are independent：$H(x, y)=F(x) G(y)$ ，the joint distri－ bution of $\left(X_{r, n}, Y_{s, n}\right)$ ：

$$
\begin{aligned}
K_{r, s}^{(n)}(x, y) & :=\operatorname{Pr}\left(X_{r, n} \leq x, Y_{s, n} \leq y\right) \\
& =\operatorname{Pr}\left(X_{r, n} \leq x\right) \operatorname{Pr}\left(Y_{s, n} \leq y\right)=F_{r, n}(x) G_{s, n}(y)
\end{aligned}
$$

When $X$ and $Y$ are not independent，the joint distribution of $\left(X_{r, n}, Y_{s, n}\right)$ ：
$K_{r, s}^{(n)}(x, y):=\operatorname{Pr}\left(X_{r, n} \leq x, Y_{s, n} \leq y\right)$
$=\operatorname{Pr}$（at least $r$ of the $X_{\ell}^{\prime} s$ are $\leq x$ ，at least $s$ of the $Y_{\ell}^{\prime} s$ are $\leq y$

$$
=\sum_{i=r}^{n} \sum_{j=s}^{n} \sum_{k} f_{k, i, j}^{(n)}(x, y)
$$

$f_{k, i, j}^{(n)}(x, y)=\frac{n!}{k!(i-k)!(j-k)!(n-i-j+k)!}(H(x, y))^{k}$

$$
\times(F(x)-H(x, y))^{i-k}(G(y)-H(x, y))^{j-k}(\bar{H}(x, y))^{n-i-j+k}
$$



$$
K_{r, s}^{(n)}(x, y)=\sum_{i=r}^{n} \sum_{j=s}^{n} \sum_{k} f_{k, i, j}^{(n)}(x, y)=K_{r, s}^{(n)}(F, G, H)
$$

Positive quadrant dependence（PQD）：

$$
H(x, y) \geq F(x) G(y) \text { for all } x, y
$$

Negative quadrant dependence（NQD）：

$$
H(x, y) \leq F(x) G(y) \quad \text { for all } x, y
$$

Dependence Structure

## Theorem 1.

For $1 \leq r, s \leq n$ ，the distribution $K_{r, s}^{(n)}$ is increasing in $H$ ．
Proof：$\frac{\partial}{\partial H} K_{r, s}^{(n)}(x, y)=n f_{r-1, s-1}^{(n-1)}(x, y) \geq 0$ ．


Figure 1：$H_{2}-H_{1}$
Figure 2：$K_{2,3}^{(3)}\left(f, G, H_{2}\right)-K_{2,3}^{(3)}\left(f, G, H_{1}\right)$

## Corollary 1.

For $1 \leq r, s \leq n$ ，the joint distribution of $\left(X_{r, n}, Y_{s, n}\right), K_{r, s}^{(n)}$ ，is PQD if $H$ is PQD，and is NQD if $H$ is NQD．

## Theoretical Applications

Baker＇s（2008）distribution：

$$
\begin{aligned}
& H_{R}^{(n)}(x, y)=\sum_{r=1}^{n} \sum_{s=1}^{n} r_{r s} F_{r, n}(x) G_{s, n}(y), \sum_{r=1}^{n} r_{s r}=\sum_{s=1}^{n} r_{s r}=\frac{1}{n}, r_{s r} \geq 0 . \\
& H_{+}^{(n)}(x, y)=\frac{1}{n} \sum_{r=1}^{n} F_{r, n}(x) G_{r, n}(y) ; H_{-}^{(n)}(x, y)=\frac{1}{n} \sum_{r=1}^{n} F_{r, n}(x) G_{n-r+1, n}(y) .
\end{aligned}
$$

Bairamov and Bayramoglu＇s（2013）distribution：
$K_{R}^{(n)}(x, y)=\sum_{r=1}^{n} \sum_{s=1}^{n} r_{r s} \operatorname{Pr}\left(X_{r, n} \leq x, Y_{s, n} \leq y\right), \sum_{r=1}^{n} r_{s r}=\sum_{s=1}^{n} r_{s r}=\frac{1}{n}, r_{s r} \geq 0$.
$K_{+}^{(n)}(x, y)=\frac{1}{n} \sum_{r=1}^{n} \operatorname{Pr}\left(X_{r, n} \leq x, Y_{r, n} \leq y\right), K_{-}^{(n)}(x, y)=\frac{1}{n} \sum_{r=1}^{n} \operatorname{Pr}\left(X_{r, n} \leq x, Y_{n-r+1, n} \leq y\right)$,

## Theorem 2.

（i）For $n \geq 1, K_{+}^{(n)} \geq H_{+}^{(n)}$ or $K_{+}^{(n)} \leq H_{+}^{(n)}$ depending on $H$ is PQD or NQD．
（ii）For $n \geq 1, K_{-}^{(n)} \geq H_{-}^{(n)}$ or $K_{-}^{(n)} \leq H_{-}^{(n)}$ depending on $H$ is PQD or NQD．
Monotonicity of $K_{+}^{(n)}(x, y)$
Fact：As $n \rightarrow \infty, H_{+}^{(n)}(x, y) \rightarrow \min \{F(x), G(y)\} \quad$（Dou et al．2013）
Problem：As $n \rightarrow \infty, H_{+}^{(n)}(x, y) \rightarrow \min \{F(x), G(y)\}$ monotonically in－ creases in $n$ ？
Theorem 3．（i）For $n \geq 2$ ，the distribution $K_{+}^{(n)}$ is of the form
$K_{+}^{(n)}=H+\frac{1}{n} \sum \sum_{i+j \leq n} \min \{i, j\}\binom{n}{i, j, n-i-j}(F-H)^{i}(G-H)^{j}(H+\bar{H})^{n-i-j}$. （ii）Let $W=(F-H)(G-H)$ and $V=H+\bar{H}$ ．Then

$$
K_{+}^{(n)}=H+\sum_{m=2}^{n} \frac{1}{m-1} \sum_{i=1}^{[m / 2]}\binom{m-1}{i, i-1, m-2 i} W^{i} V^{m-2 i}, \quad n \geq 2
$$

where $[a]$ is the largest integer less than or equal to $a$ ．Equivalently，

$$
K_{+}^{(n)}-K_{+}^{(n-1)}=\frac{1}{n-1} \sum_{i=1}^{[n / 2]}\binom{n-1}{i, i-1, n-2 i} W^{i} V^{n-2 i} \geq 0, \quad n \geq 2
$$

## References

－Bairamov，I．and Bayramoglu，K．（2013）．From the Huang－Kotz FGM distribution to Baker＇s bivariate distribution，Journal of Multivariate Analysis，113，106－115．
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